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Research Article

Hydroxychloroquine for Pre-Exposure Prophylaxis for Sars-Cov-2

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Abstract

SARS-CoV-2 infection has a high transmission level. At the present time there is not a specific treatment approved but it is known that, in vitro, chloroquine and hydroxychloroquine can inhibit the coronavirus.

Objective: verifying if patients with autoimmune diseases that are on treatment with HCQ have less incidence and severity on COVID-19.

Material and methods: this is a retrospective cohort study. The exposed cohort was formed by individuals with autoimmune diseases with HCQ treatment. The control cohort was randomly selected using the Health Card database. To deal with confounding variables and evaluate the effect of HCO on the incidence and severity of SARS-CoV-2 infection, propensity score matching was used. Risk difference and paired percentage difference between exposed and non-exposed groups was estimated.

Results: 919 individuals formed the exposed cohort and 1351 the control cohort. After matching, there were 690 patients on each group. During the time of the study, in the exposed group there were 42 (6.1%) individuals with suspected COVID-19, 12(1.7%) with confirmed COVID-19 and 3(0.4%) were hospitalized. In the control group there were 30(4.3%) individuals with suspected COVID-19, 13(1.9%) with confirmed COVID-19 and 2(0.3%) were hospitalized. The risk difference between each cohort was: 0.017(-0.05-0.04) for suspected COVID-19; -0.014(-0.015-0.012) for confirmed COVID-19 and 0.001(-0.007-0.007) for hospitalized patients. There were not significant differences.

Conclusion: there is no difference neither on the incidence nor on the severity of COVID-19 between patients with autoimmune diseases with HCQ treatment and patients that do not take HCQ.

Strengths and limitations of this study

Strenghts: Ii provides evidence on pre-exposure prophylaxis of HCQ against COVID-19 for which there is still no evidence in clinical trials.

The study on the pre-exposure prophylaxis of HCQ was carried out during the period of maximum circulation of SARS-COV-2 in Spain. Therefore, both cohorts had a high probability of exposure to the virus, which increases the power of the study.

The dose of HCQ taken by most of the study subjects (400 mg per day) is similar to that used in several clinical trials that are underway as pre- and post-exposure prophylaxis of COVID-19.

Limitations: The population of the exposed group suffers entirely from rheumatic diseases. This condition could act as a confounding factor, making it difficult to generalize the results to the entire population.

Propensity score matching resulted in a reduction in sample size. Of the 900 initial couples, only 690 survive.

Given the shortage of diagnostic tests during the study period, many of the possible COVID-19s were left unconfirmed, limiting the power of the study.

Keywords: Hydroxychloroquine, Pre-exposure prophylaxis, COVID-19, SARS-CoV-2

Introduction

The big morbi-mortality plus the lack of treatments and specific vaccines against SARS-CoV-2 virus, made the scientific community use drugs that had already been used to fight other diseases.

Chloroquine (CQ) and hydroxychloroquine (HCQ) used in vitro [1,2] have shown effectiveness against other viruses, including the responsible of the previous severe acute respiratory syndrome SARS-CoV coronavirus outbreak. This findings plus the wide experience using these drugs and its low price, made them good candidates to be used for prophylaxis and treatment for COVID-19.

In vitro trials in China show that CQ and HCQ inhibit the growth of SARS-CoV-2, showing a good antiviral activity for pre-exposure prophylaxis and treatment. HCQ has a stronger effect as a pre-exposure prophylaxis treatment [3]. Other in vitro trials suggest different dosage for both prophylaxis and treatment for SARS-CoV-2 [4].

Small trials show effectiveness using HCQ both on its own [5] and in combination with macrolides [6,7] to treat COVID-19. However, more recent evidence demonstrates the contrary [8-10].

The first trial speculating on a post-exposure prophylactic role of HCQ took place in a hospital in South Korea [11]. There was not a control group. HCQ was given to 211 people (189 patients and 22 hospital workers) with a negative result in the Polymerase Chain Reaction (PCR) diagnose, that had previously been in touch with a COVID-19 patient. After a 14 days quarantine, 97% of the sub-

jects kept a negative PCR.

Some clinical trials [12] and some publications in the literature [13] say that patients treated with HCQ were less affected or not affected at all by COVID-19. The potential validity of these findings was used by some politicians to make statements about the benefits of antimalarials in the prophylaxis of SARS-CoV-2 [14]. Later publications in the COVID-19 Global Rheumatology Alliance express the opposite [15,16].

Whilst ongoing clinical trials on the prophylaxis of COVID-19 reveal results, we present this study with the aim of evaluating if patients who are on a chronic treatment with antimalarials have less incidence of infection with SARS-CoV-2 and/or a less severe disease, than patients who do not take antimalarials.

Materials and Methods

Study type: Retrospective cohort study.

Data source: The *exposed cohort* is formed by all the patients on a chronic treatment with HCQ in León (Spain). Mostly due to autoimmune diseases: systemic lupus erythematosus (LES) 43%, rheumatoid arthritis (RA) 28% and other rheumatic diseases 29%. This cohort was elaborated with the CONCILYA database, that holds information about the invoicing of pharmaceutical prescriptions. Patients who had been prescribed with CQ or HCQ during December 2019, January and February 2020 were selected using the Patient Identification Number (PIN). We assume that considering this range of time, we selected all the patients that take antimalarials on a chronic base. Data was firstly preprocessed, removing 37 patients who had only withdrawn one package of either CQ or HCQ from the pharmacy in March, verifying that they were not diagnosed with a rheumatic disease. We suppose they had withdrawn it as a possible prevention for COVID-19. Another 42 patients were removed as they did not belong to the study area.

The *non-exposed cohort* was selected randomly samplingfrom the Individual Health Card (IHC) database of the population from the target area. They were matched by sex, due to the difference in the prevalence of rheumatic diseases between men and women, and by a five-years age range.

Sample size and power of the study

The size of the exposed cohort, 919, is given by the number of

patients with autoimmune diseases (LES and RA mostly) that take CQ or HCQ. 1351 subjects of the non-exposed cohort were selected in order to get a 95% confidence. In 900 pairs of subjects we would detect risk differences of 95% or more with an 80% power.

Study variables

Exposure variables: taking or not CQ or HCQ. It was classified as a dichotomous variable (yes/no).

Output variables: having COVID-19 disease during the months of maximum impact of the pandemic (March and April 2020). This variable was studied in two different levels: possible disease or confirmed disease. Possible corresponds to patients with COVID-19-like symptoms such as non-severe acute respiratory infection (fever, cough, dyspnea, myalgia, shivers), ageusia or anosmia, during the time of the study, with no diagnostic tests. Confirmed, corresponds to having a positive Polymerase Chain Reaction (PCR) test for SARS-CoV-2 in a nasopharyngeal swab or having COVID-19-like symptoms and a positive COVID-19 Rapid Antibody Test (IgM and/or IgG). Being hospitalized was also evaluated to measure the severity of the disease.

Other variables: sociodemographic data: age, sex, rural or urban area, sanitary professional, workers at elderly homes; disease for which they take antimalarials; coagulopathy; diseases and drugs that could affect on the morbimortality of COVID-19: obesity, smoker, hypertension (HT), lung disease, cardiovascular disease (CVD); patients on treatment with: corticosteroids, angiotensin-converting-enzyme inhibitors (ACE inhibitor) / angiotensin II receptor blockers (ARBs), anticoagulant medication, vitamin D, non-steroidal anti-inflammatory (NSAIDs).

Fieldwork: Using the PIN, we had access to each patient personal history database in MEDORA (used in Primary Care) and JIME-NA (used in the hospital). It was checked on every patient if they had suffered COVID-19, which symptoms they had and if they had been hospitalized. Each patient was contacted via telephone to obtain variables that were not present on their personal history, such as previous exposure to the virus. May 1st 2020 was considered the study closure date.

Statistical analysis

A descriptive study of each variable was carried out using frequen-

cy distributions for qualitative variables, means and standard deviations for quantitative ones.

This is an observational study. Thus, in order to estimate causal effects of treatment with CQ or HCQ on the different outcomes, it is necessary to control for confounding. For this purpose we use propensity score mating (PSM) [17]. PSM first estimates the propensity score (i.e. the conditional probability of receiving treatment given the confounding variables) for all individuals. Subsequently, individuals of both groups (treatment and control) are matched by propensity scores. As PSM cannot deal with missing data, an imputation process was previously carried out using multiple imputation by chained equations (MICE) [18]. A sensitivity analysis of the conclusions of the study against different imputations of missing values is presented in the Appendix.

The propensity score was estimated using a logistic regression model. The confounding variables included in the model were those that, based on the results of the raw analyses and the bibliography consulted [19], could behave as confounding factors. Specifically, they were following ones: age, sex, smoker, HT, Diabetes, CVD, Lung disease, anticoagulant medication, corticosteroids, vitamin D, NSAIDs, and possible exposure to COVID. A oneto-three matching was performed using "closest neighbors" technique [20]. A caliper of 0.2 was used (that is, pairs whose distance between propensity scores is greater than 0.2 standard deviations are not accepted). After matching, to ensure balance, differences between treatment and control groups in the all the covariates are checked using standardized mean deviations (SMDs). In the appendix, the conclusions of the study are shown for one-to-two and one-to-four matching.

Finally, after performing PSM, causal effects are estimated in the paired sample. To that end, the risk difference between treatment and control groups was studied for all three outcome variables (using a paired Student test), as well as the difference in paired percentages (through a McNemar test).

All the statistical calculations of the present study were performed using the R statistical software, version 3.6.3. The imputation of missing values was made with the R [21] "mice" library and the propensity score "matching" with this same software [22].

Results

Descriptive analysis and PSM matching

There were 919 patients with rheumatic diseases on treatment with antimalarials. 55% took 400mg of HCQ per day; 44%, 200mg or less and only 1% took CQ. 85% were on treatment for more than one year.

4% of the patients in the non-treated group and 5.4% of those in the treatment group were diagnosed with possible COVID-19. On 1.6% of each group the diagnosis was confirmed. Table 1 shows the characteristics of the treatment and control groups, before and after matching.

As shown by the SMD, before PSM there were great differencesbetween exposed and control groups in covariables such as HT, CVD, lung disease, treatment with corticosteroids, vitamin D, anticoagulants or NSAIDs. This was expectable, as patients with autoimmune diseases (who were on treatment with antimalarials) were more likely to suffer HT or lung diseases. In addition, there is also a strong correlationbetween taking antimalarials and vitamin D or corticosteroids since this drugs are also part of the treatment for autoimmune diseases.

These variables may act as confounders. To achieve an adequate balance in treatment and control groups and be able to extract causal conclusions we use PSM. As shown in Table 1, after PSM, 690 patients survive within each cohort. Now, the differences in the covariates between the two groups are much smaller, as shown by the SMD. Thus, the matched sample can be considered to be balanced.

Causal effects estimation

The following results were obtained using matched data: in the exposed cohort there were 42 (6.1%) individuals with suspected COVID-19; 12 (1.7%) individuals with confirmed COVID-19 and 3 (0.4%) individuals were hospitalized. In the non-exposed cohort there were 30 (4.3%) individuals with suspected COVID-19; 13 (1.9%) individuals with confirmed COVID-19 and 2 (0.3%) individuals were hospitalized.

The risk difference between each cohort, as shown in Table 2, was: 0.017 (-0.05-0.04) for suspected COVID-19; -0.014 (-0.015-0.012)

Table 1: Individuals characteristics depending on whether if they took HCQ before or after the matching with PSM.									
		Original sample	2	Matchingsampleafter PSM					
	Nottreated- with HCQ	Treatedwith HCQ	SMD	Nottreated- with HCQ	Treatedwith HCQ	SMD			
n	1351	919		690	690				
Age (mean(SD))	58.96 (15.91)	57.28 (15.03)	0,109	55.66 (16.76)	57.40 (14.19)	0,112			
Women, n (%)	1055 (78.1)	706 (76.8)	0,03	530 (76.8)	522 (75.7)	0,027			
Profession, n (%)			0,049						
Healthcareprofessional	23 (1.7)	21 (2.3)							
Elderlyhomes	33 (2.4)	26 (2.8)							
Others	1295 (95.9)	872 (94.9)							
Smoker, n (%)	227 (16.8)	164 (17.8)	0,028	130 (18.8)	147 (21.3)	0,062			
Urbanarea, n (%)	756 (56.0)	499 (54.3)	0,033						
Obesity, n (%)	251 (18.6)	209 (22.7)	0,103	130 (18.8)	144 (20.9)	0,051			
HT, n (%)	369 (27.3)	322 (35.0)	0,167	198 (28.7)	213 (30.9)	0,048			
Diabetes, n (%)	130 (9.6)	93 (10.1)	0,017	55 (8.0)	63 (9.1)	0,041			
CVD, n (%)	64 (4.7)	100 (10.9)	0,231	43 (6.2)	63 (9.1)	0,109			
Lungdisease, n (%)	24 (1.8)	39 (4.2)	0,145	16 (2.3)	24 (3.5)	0,069			
Take ACE inhibitors, n (%)	143 (10.6)	99 (10.8)	0,006						
TakeARBs, n (%)	144 (10.7)	139 (15.1)	0,134						
Takeanticoagulants, n (%)	58 (4.3)	112 (12.2)	0,29	37 (5.4)	45 (6.5)	0,049			
Takecorticosteroids n (%)	27 (2.0)	346 (37.6)	1	25 (3.6)	39 (5.7)	0,097			
Take vitamin D, n (%)	144 (10.7)	644 (70.1)	1,522	140 (20.3)	138 (20.0)	0,007			
TakeNSAIDs, n (%)	113 (8.4)	234 (25.5)	0,468	100 (14.5)	108 (15.7)	0,032			
Drinktonic, n (%)	22 (1.6)	46 (5.0)	0,189						
Work COVID-19 exposure, n (%)	50 (3.7)	52 (5.7)	0,093	36 (5.2)	30 (4.3)	0,041			
Familiar COVID-19 exposure, n (%)	43 (3.2)	26 (2.8)	0,021	19 (2.8)	30 (4.3)	0,086			
Suspected COVID-19, n (%)	54 (4)	50 (5,4)	0,021	30 (4.3)	42 (6.1)	0,078			
Confirmed COVID-19, n (%)	22 (1.6)	15 (1.6)	<0.001	13 (1.9)	12 (1.7)	0,011			
Hospitalized for COVID-19, n (%)	5 (0.4) *	6 (0.7)	0.04	2 (0.3)*	3 (0.4)	0,024			

SD: standard desviation; SMD: standarized mean deviation. HT: hypertension, ACE inhibitors: angiotensin-converting-enzyme inhibitors, ARBs: angiotensine II receptor blockets, NSAIDs: non-steroidal anti-inflamatory.

 $\ensuremath{^*\!:}$ one patient passed on each group.

lacksquare

for confirmed COVID-19 and 0.001 (-0.007-0.007) for hospitalized patients. There were not significant differences between each cohort on neither of the three variables.

Discussion

Our results show that being on a chronic treatment with either CQ or HCQ has neither benefit on the pre-exposure prophylaxis for SARS-CoV-2 nor on the avoidance of being hospitalized, as a subrogate variable for severity of the infection.

Age, sex, rural or urban area and being an active smoker were balanced on the original sample in both groups. On the contrary, obesity, CVD, HT and lung disease were more frequent on the group formed by individuals who were taking HCQ due to a rheumatic disease. These illness association with autoimmune diseases [23] could have affected the final results [24]. These same comorbidities plus advanced age, male sex and diabetes were associated with a worse outcome on the SARS-CoV-2 infection [25,26]. A great number of individuals on the exposed cohort were on treatment with corticosteroids, immunosuppressive drugs or biological therapies. There is a controversy in the conclusions of different studies about COVID-19 severity in patients with autoimmune diseases who are on treatment with this kind of drugs [27-30]. All these facts could make us believe that the results of our trial would show a protective effect from the antimalarial drugs. However, after matching, there is no significative difference between both groups neither in the number of hospitalized patients nor in the mortality of COVID-19.

Our results are in line with other studies that reveal a lack of evidence on the efficiency of HCQ for prophylaxis and treatment for COVID-19 [31]. A meta-analysis shows poor evidence on the efficiency of CQ and HCQ on the prevention of COVID-19 [32]. Likewise, Boulmare DR et al on a clinical trial in the USA and Canada [33], conclude that HCQ shows no efficiency on the post-exposure prevention neither for suspected nor for confirmed COVID-19. In a Spanish trial on autoimmune inflammatory diseases, Macias, J. et al [34], reveal a similar result. They also observe no differences with placebo neither on the treatment for suspected or confirmed COVID-19 nor on the hospitalized individuals.

400mg of HCQ per day was the dosage used by most of our individuals. The same dosage is being used in other ongoing trials for pre-exposure and post-exposure prophylaxis of COVID-19, shown in clinicaltrials.gov (NCT 04333225, NCT 04331834). This dosage is bigger than the one proposed in other in vitro studies [6] for pre-exposure prophylaxis.

There are several limitations in our study. This is a retrospective study that analyzes real-life data of two groups, treatment and control, whose basal characteristics were not comparable. To overcome this limitation PSM was needed. The size of our sample was reduced because it was not possible to match some individuals.

The estimated number of matches needed to achieve a correct output for the hypothesis tests was 900. Neverthless, after matching, sample size was reduced to 690 patients per group. However, we end up with two comparable groups where the only difference was treatment status before exposure.

In Spain most COVID-19 cases were not severe and therefore were well managed by General Practitioners from Primary Care. At the beginning of the pandemic there were not enough diagnostic tests, hence they were mostly used for hospitalized patients. This is the reason why a considerable number of suspected COVID-19

	Suspected COVID		Confirm	Confirmed COVID		Hospitalized for COVID	
	Yes	No	Yes	No	Yes	No	
Treatmentwith HCQ	42	648	12	678	3	687	
Nottreatmentwith HCQ	30	660	13	677	2	688	
Riskdifference (I.C. 95%)	0.017 (-0.005;0.04)		-0.0014 (-0.015;0.012)		0,001 (-0.007;0.007)		
McNemar'schisquare	1.8; p-value=0.17		0.001; p-value=0,662		0.00; p-value=1		

Table 2: HCQ effect on having suspected COVID-19, confirmed COVID-19 and being hospitalized for COVID-19, after PSM.

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patients were not confirmed. Of 72 suspected individuals only 25 were confirmed with SARS-CoV-2 infection. The 47 remaining, exposed and non-exposed, were not confirmed. If we would have been able to perform diagnostic tests maybe the confirmed patients would have inclined the balance to either the exposed group or to the non-exposed group.

Having used a pragmatic approach made it easier to recruit the exposed cohort but it made it more difficult to apply these results to general population. All the individuals in the exposed group had rheumatic diseases. This could act as a confounding factor because it presence of rheumatic diseases may be associated with some or the wholesome of the effect variables, this complicating the generalization of the results. Nevertheless, in the hypothetical case where having a rheumatic disease did not affect the outcome variables, the conclusions in this study could be applied to general population. This could confirm the lack of positive benefit of HCQ neither for pre-exposure prophylaxis nor for diminishing severity of COVID-19.

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Supplementary Appendix

This appendix has been provided by the authors to give readers additional information about their work.

Sensitivity analysis

Using MICE[18], 3 different missing value imputations in addition to the one presented in the study were performed. As shown in Tables A1, A2 and A3, the main conclusions do not change.

For this study, a one-to-three propensity score matching [17] was performed. Tables A4 and A5 show the main results under one-to-two and one-to-four matching, respectively. On the one-to-two matching case, the final sample size was 530, while on the one-to-four it was 836. The conclusions remain unchanged in both cases.

Table A1: HCQ effect on having suspected COVID-19, confirmed COVID-19 and being hospitalized for COVID-19, after matching with PSM. Imputation 1.									
	Suspected COVID		Confirmed COVID		Hospitalized for COVID				
	Yes	No	Yes	No	Yes	No			
Treatment with HCQ	42	651	12	681	3	690			
Not treatment with HCQ	30	663	14	679	3	690			
Risk difference (I.C.95%)	0.017 (-0.006;0.04)		0,0029 (-0.017;0.012)		0,00 (-0.007;0.007)				
McNemar's chi square	1.7; p-value=0.19		0.038; p-value=0,845		0.00; p-value=1				



Table A2: HCQ effect on having suspected COVID-19, confirmed COVID-19 and being hospitalized for COVID-19, after matching with PSM. Imputation 2.

	Suspected COVID		Confirmed COVID		Hospitalized for COVID	
	Yes	No	Yes	No	Yes	No
Treatment with HCQ	42	651	12	681	3	690
Not treatment with HCQ	35	668	15	678	4	689
Risk difference (I.C. 95%)	0.010 (-0.014;0.034)		0.0043 (-0.019;0.001)		-0.001 (-0.008;0.005)	
McNemar's chi square	0.5; p-value=0.47		0.16; p-value=0.689		0.00; p-value=1	

Table A3: HCQ effect on having suspected COVID-19, confirmed COVID-19 and being hospitalized for COVID-19, after matching with PSM. Imputation 3.

	Suspected COVID		Confirmed COVID		Hospitalized for COVID	
	Yes	No	Yes	No	Yes	No
Treatment with HCQ	39	651	9	681	0	690
Not treatment with HCQ	31	659	15	675	3	687
Risk difference	0.012		-0,0087		-0.004	
(I.C. 95%)	(-0.012;0.035)		(-0,023;0,005)		(-0.009;0.0005)	
McNemar's chi square	0.74;		1.04;		1.33;	
	p-value=0.39		p-value=0.307		p-value=0.25	

Table A4: HCQ effect on having suspected COVID-19, confirmed COVID-19 and being hospitalized for COVID-19, after matching one-to-two with PSM.

	Suspected COV	/ID	Confirmed Co	OVID	Hospitalized for COVID	
	Yes	No	Yes	No	Yes	No
Treatment with HCQ	30	500	8	522	2	528
Not treatment with HCQ	22	508	11	519	2	528
Risk difference (I.C. 95%)	0.015 (-0.01;0.04)		0.0057 (-0.022;0.011)		0.00 (-0.007;0.007)	
McNemar's chi square	1.1; p-value=0.302		0.211; p-value=0.646		0.00; p-value=1	

Table A5: HCQ effect on having suspected COVID-19, confirmed COVID-19 and being hospitalized for COVID-19, after matching one-to-four with PSM.

	Suspected COVID		Confirmed COVID		Hospitalized for COVID	
	Yes	No	Yes	No	Yes	No
Treatment with HCQ	48	788	12	824	0	836
Not treatment with HCQ	35	801	15	821	2	834
Risk difference (I.C. 95%)	0.015 (-0.006;0,04)		-0.0036 (-0.016;0.009)		0.002 (-0.006;0.0009)	
McNemar's chi square	0.8; p-value=0.18		0.15; p-value=0.700		0.5; p-value=0.48	

